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Why estimate production functions?

- Primitive component of economic model
- Gives estimate of firm productivity useful for understanding economic growth
 - Stylized facts to inform theory, e.g. Foster, Haltiwanger, and Krizan (2001)
 - Effect of deregulation, e.g. Olley and Pakes (1996)
 - Growth within old firms vs from entry of new firms, e.g. Foster, Haltiwanger, and Krizan (2006)
 - Effect of trade liberalization, e.g. Amiti and Konings (2007)
 - Effect of FDI Javorcik (2004)

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General references:

- Aguirregabiria (2021) chapter 3
- De Loecker and Syverson (2021)
- Ackerberg et al. (2007) section 2
- Van Beveren (2012)

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• Cobb Douglas production

$$Y_{it} = A_{it} K_{it}^{\beta_k} L_{it}^{\beta_\ell}$$

In logs,

$$y_{it} = \beta_k k_{it} + \beta_\ell I_{it} + \omega_{it} + \epsilon_{it}$$
with $\log A_{it} = \omega_{it} + \epsilon_{it}$, ω_{it} known to firm, ϵ_{it} not

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Empirical Challenges

- ① Simultaneity: if firm has information about $\log A_{it}$ when choosing inputs, then inputs correlated with $\log A_{it}$
- Selection: firms with low productivity will exit sooner
- 3 Others: measurement error, specification

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Simultaneity

- Firm's inputs will be correlated with firm's knowledge of productivity
- E.g. output price p, wage w, choosing L given K

$$\max_{l} p E[A] K_k^{\beta} L_{\ell}^{\beta} - wL$$

implies

$$L = \left(\frac{p}{w}\beta_{\ell} \mathsf{E}[A]K^{\beta_k}\right)^{\frac{1}{1-\beta_{\ell}}}$$

or in logs,

$$I = \frac{1}{1 - \beta_{\ell}} \log \left(\frac{p}{w} \beta_{\ell} \right) + \frac{\beta_{k}}{1 - \beta_{\ell}} k + \frac{1}{1 - \beta_{\ell}} \log \left(\mathbb{E}[A] \right)$$

- I correlated with productivity through correlation of log A and log(E[A])
- Exercise: calculate bias of $\hat{\beta}_{\ell}^{OLS}$ (with some further assumptions)

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Simultaneity solutions

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Instrumental variables

- Instrument must be
 - Correlated with k and l
 - Uncorrelated with $\omega + \epsilon$
- Possible instrument: input prices
 - Correlated with *k*, *l* through first-order condition
 - Uncorrelated with ω if input market competitive
- Other possible instruments: output prices (more often endogenous), input supply or output demand shifter (hard to find)

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Problems with input prices as IV

- Not available in many data sets
- Average input price of firm could reflect quality as well as price differences
- Need variation across observations
 - If firms use homogeneous inputs, and operate in the same output and input markets, we should not expect to find any significant cross-sectional variation in input prices
 - If firms have different input markets, maybe variation in input prices, but different prices could be due to different average productivity across input markets
 - Variation across time is potentially endogenous because could be driven by time series variation in average productivity

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Fixed effects

- Have panel data, so should consider fixed effects
- FE consistent if:
 - $\mathbf{1} \ \omega_{it} = \eta_i + \delta_t + \omega_{it}^*$
 - 2 ω_{it}^* uncorrelated with l_{it} and k_{it} , e.g. ω_{it}^* only known to firm after choosing inputs
 - 3 ω_{it}^* not serially correlated and is strictly exogenous
- Problems:
 - Fixed productivity a strong assumption
 - Estimates often small in practice
 - Worsens measurement error problems

$$\mathsf{Bias}(\hat{eta}_k^{\mathsf{FE}}) pprox - \frac{eta_k \mathsf{Var}(\Delta \epsilon)}{\mathsf{Var}(\Delta k) + \mathsf{Var}(\Delta \epsilon)}$$

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Dynamic panel: motivation 1

- General idea: relax fixed effects assumption, but still exploit panel
- Collinearity problem: Cobb-Douglas production, flexible labor and capital implies log labor and log capital are linear functions of prices and productivity (Bond and Söderbom (2005))
- If observed labor and capital are not collinear then there must be something unobserved that varies across firms (e.g. prices), but that could invalidate monotonicity assumption of control function

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Dynamic panel: moment conditions

- See Blundell and Bond (2000)
- Assume $\omega_{it} = \gamma_t + \eta_i + \nu_{it}$ with $\nu_{it} = \rho \nu_{i,t-1} + e_{it}$, so

$$y_{it} = \beta_{\ell}I_{it} + \beta_{k}k_{it} + \gamma_{t} + \eta_{i} + \nu_{it} + \epsilon_{it}$$

subtract $\rho y_{i,t-1}$ and rearrange to get

$$y_{it} = \rho y_{i,t-1} + \beta_{\ell} (I_{it} - \rho I_{i,t-1}) + \beta_{k} (k_{it} - \rho k_{i,t-1}) +$$

$$+ \gamma_{t} - \rho \gamma_{t-1} + \underbrace{\eta_{i} (1 - \rho)}_{=\eta_{i}^{*}} + \underbrace{e_{it} + \epsilon_{it} - \rho \epsilon_{i,t-1}}_{=w_{it}}$$

- Moment conditions:
 - Difference: $E[x_{i,t-s}\Delta w_{it}] = 0$ where x = (l, k, y)
 - Level: $E[\Delta x_{i,t-s}(\eta_i^* + w_{it})] = 0$

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Dynamic panel: economic model 1

Adjustment costs

$$\begin{split} V(K_{t-1}, L_{t-1}) &= \max_{I_t, K_t, H_t, L_t} P_t F_t(K_t, L_t) - P_t^K (I_t + G_t(I_t, K_{t-1})) - \\ &- W_t \left(L_t + C_t(H_t, L_{t-1}) \right) + \\ \psi \mathbb{E} \left[V(K_t, L_t) \middle| \mathcal{I}_t \right] \\ \text{s.t. } K_t &= (1 - \delta_k) K_{t-1} + I_t \\ L_t &= (1 - \delta_l) L_{t-1} + H_t \end{split}$$

Implies

$$P_{t} \frac{\partial F_{t}}{\partial L_{t}} - W_{t} \frac{\partial C_{t}}{\partial L_{t}} = W_{t} + \lambda_{t}^{L} \left(1 - (1 - \delta_{l}) \psi \mathbb{E} \left[\frac{\lambda_{t+1}^{L}}{\lambda_{t}^{L}} | \mathcal{I}_{t} \right] \right)$$

$$P_{t} \frac{\partial F_{t}}{\partial K_{t}} - P_{t}^{K} \frac{\partial G_{t}}{\partial K_{t}} = \lambda_{t}^{K} \left(1 - (1 - \delta_{k}) \psi \mathbb{E} \left[\frac{\lambda_{t+1}^{K}}{\lambda_{t}^{K}} | \mathcal{I}_{t} \right] \right)$$

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Dynamic panel: economic model 2

- Current productivity shifts $\frac{\partial F_t}{\partial L_t}$ and (if correlated with future) the shadow value of future labor $\mathbb{E}\left[\frac{\lambda_{t+1}^L}{\lambda_t^L}|\mathcal{I}_t\right]$
- Past labor correlated with current because of adjustment costs

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Dynamic panel data: problems

Problems:

- Sometimes imprecise (especially if only use difference moment conditions)
- Differencing worsens measurement error
- Weak instrument issues if only use difference moment conditions but levels stronger (see Blundell and Bond (2000))
 - Level moments require stronger stationarity assumption
 - η_i uncorrelated with Δx_{it}

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Control functions

- From Olley and Pakes (1996) (OP)
- Control function: function of data conditional on which endogeneity problem solved
 - E.g. usual 2SLS $y = x\beta + \epsilon$, $x = z\pi + v$, control function is to estimate residual of reduced form, \hat{v} and then regress y on x and \hat{v} . \hat{v} is the control function
- Main idea: model choice of inputs to find a control function

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OP assumptions

$$y_{it} = \beta_k k_{it} + \beta_\ell I_{it} + \omega_{it} + \epsilon_{it}$$

1 ω_{it} follows exogenous first order Markov process,

$$p(\omega_{it+1}|\mathcal{I}_{it}) = p(\omega_{it+1}|\omega_{it})$$

2 Capital at t determined by investment at time t-1,

$$k_t = (1 - \delta)k_{it-1} + i_{it-1}$$

$$i_{it} = I_t(k_{it}, \omega_{it}),$$

and is strictly increasing in ω_{it}

4 Labor variable and non-dynamic, i.e. chosen each *t*, current choice has no effect on future (can be relaxed)

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OP estimation of β_{ℓ}

• Invertible investment implies $\omega_{it} = I_t^{-1}(k_{it}, i_{it})$

$$y_{it} = \beta_k k_{it} + \beta_\ell I_{it} + I_t^{-1}(k_{it}, I_{it}) + \epsilon_{it}$$

= $\beta_\ell I_{it} + f_t(k_{it}, I_{it}) + \epsilon_{it}$

- Partially linear model
 - Estimate by e.g. regress y_{it} on l_{it} and series functions of t, k_{it}, i_{it}
 - Gives $\hat{\beta}_l$, $\hat{f}_{it} = \hat{f}_t(k_{it}, i_{it})$

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OP estimation of β_k

- Note: $\hat{f}_t(k_{it}, i_{it}) = \hat{\omega}_{it} + \beta_k k_{it}$
- By assumptions, $\omega_{it} = \mathbb{E}[\omega_{it}|\omega_{it-1}] + \xi_{it} = g(\omega_{it-1}) + \xi_{it}$ with $\mathbb{E}[\xi_{it}|k_{it}] = 0$
- Use $E[\xi_{it}|k_{it}] = 0$ as moment to estimate β_k .
 - OP: write production function as

$$y_{it} - \beta_{\ell} I_{it} = \beta_k k_{it} + g(\omega_{it-1}) + \xi_{it} + \epsilon_{it}$$

$$= \beta_k k_{it} + g(f_{it-1} - \beta_k k_{it-1}) + \xi_{it} + \epsilon_{it}$$

Use $\hat{\beta}_l$ and \hat{f}_{it} in equation above and estimate $\hat{\beta}_k$ by e.g. semi-parametric nonlinear least squares

• Ackerberg, Caves, and Frazer (2015): use $\mathbb{E}\left[\hat{\zeta}_{it}(\beta_k)k_{it}\right]=0$

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Control functions

Dynamic panel vs control function

- Both derive moment conditions from assumptions about timing and information set of firm
- Dealing with ω
 - Dynamic panel: AR(1) assumption allows quasi-differencing
 - Control function: makes ω estimable function of observables
- Dynamic panel allows fixed effects, does not make assumptions about input demand
- Control function allows more flexible process for ω_{it}

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- Olley and Pakes (1996): productivity in telecom after deregulation
- Söderbom, Teal, and Harding (2006): productivity and exit of African manufacturing firms, uses IV
- Levinsohn and Petrin (2003): compare estimation methods using Chilean data
- Javorcik (2004): FDI and productivity, uses OP
- Amiti and Konings (2007): trade liberalization in Indonesia, uses OP
- Aw, Chen, and Roberts (2001): productivity differentials and firm turnover in Taiwan
 - Kortum and Lerner (2000): venture capital and innovation

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Selection

- Let $d_{it} = 1$ if firm in sample.
 - Standard conditions imply $d = 1\{\omega \ge \omega^*(k)\}$
- Messes up moment conditions
 - All estimators based on $E[\omega_{it}Something] = 0$, observed data really use $E[\omega_{it}Something|d_{it} = 1]$
 - E.g. OLS okay if $E[\omega_{it}|I_{it},k_{it}]=0$, but even then,

$$E[\omega_{it}|I_{it}, k_{it}, d_{it} = 1] = E[\omega_{it}|I_{it}, k_{it}, \omega_{it} \ge \omega^*(k_{it})]$$
$$= \lambda(k_{it}) \ne 0$$

• Selection bias negative, larger for capital than labor

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Selection in OP model

- Estimate β_{ℓ} as above
- Write $d_{it} = 1\{\xi_{it} \leq \omega^*(k_{it}) \rho(f_{i,t-1} \beta_k k_{it-1}) = h(k_{it}, f_{it-1}, k_{it-1})\}$
- Propensity score $P_{it} \equiv \mathbb{E}[d_{it}|k_{it},f_{it-1},k_{it-1}]$
- Similar to before estimate β_k , from

$$y_{it} - \beta_{\ell} I_{it} = \beta_k k_{it} + \tilde{g} (f_{it-1} - \beta_k k_{it-1}, P_{it}) + \xi_{it} + \epsilon_{it}$$

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